# FOUNDATIONS OF MARTINGALE THEORY AND STOCHASTIC CALCULUS FROM A FINANCE PERSPECTIVE

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# 1. INTRODUCTION

The language of mathematical Finance allows to express many results of martingale theory via trading arguments, which makes it somehow easier to appreciate their contents. Just to provide one illustrative example: let  $X = (X_n)_{n\geq 0}$  be a martingale and let  $N_a^b(t)$  count the upcrossings over the interval [a, b], then we can find a predictable process  $V = (V_n)_{n\geq 0}$  such that

$$(b-a)N_a^b(t) \le (a-X_t)^+ + (V \bullet X)_t,$$

where the right hand side's stochastic integral corresponds precisely to the cumulative gains and losses of a buy below a and sell above b strategy plus an option payoff adding for the risk of buying low and loosing even more. Immediately Doob's upcrossing inequality follows by taking expectations. We shall somehow focus on such types of arguments in the sequel, however, pushing them as far as possible. Still most of the following proofs are adapted from Olav Kallenberg's excellent book [3] or from the excellent lecture notes [6], maybe with a bit different wording in each case. For stochastic integration the reading of [6] is recommended.

Another important aspect is the two-sided character of many arguments, which leads, e.g., to reverse martingale results.

# 2. Filtrations, Stopping times, and all that

Given a probability space  $(\Omega, \mathcal{F}, P)$ , a *filtration* is an increasing family of sub- $\sigma$ -algebras  $(\mathcal{F}_t)_{t \in T}$  for a given index set  $T \subset \mathbb{R} \cup \{\pm \infty\}$ .

We shall often assume the "usual conditions" on a filtered probability space, i.e. that a filtration is right continuous and complete, but we first name the properties separately:

- (1) A filtration is called *complete* if each  $\mathcal{F}_t$  contains all *P*-null sets from  $\mathcal{F}$  for  $t \in T$ . In particular every  $\sigma$ -algebra  $\mathcal{F}_t$  is complete with respect to *P*.
- (2) A filtration is called *right continuous* if  $\mathcal{F}_t = \mathcal{F}_{t+} := \bigcap_{\epsilon > 0} \mathcal{F}_{t+\epsilon}$  for all  $t \in T$ , where  $\mathcal{F}_{t+\epsilon} := \bigcap_{u \in T, u \ge t+\epsilon} \mathcal{F}_u$  for  $\epsilon > 0$ .

Apparently every filtration  $(\mathcal{F}_t)_{t\in T}$  has a smallest right continuous and complete filtration  $(\mathcal{G}_t)_{t\in T}$  extending in the sense that  $\mathcal{F}_t \subset \mathcal{G}_t$  for  $t \in T$ , namely

$$\mathcal{G}_t = \overline{\mathcal{F}_{t+}} = \overline{\mathcal{F}_t}_+ \,.$$

It is called the *augmented filtration*.

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Notice that usual conditions are necessary if one expects P-nullsets to be part of  $\mathcal{F}_0$  to guarantee the process to be adapted, which appears in constructions like regularization procedures. For most purposes right continuous filtrations are enough to obtain regularized processes (e.g. with cadlag trajectories) outside a P-nullset.

**Definition 2.1.**  $A \ T \cup \{\infty\}$ -valued random variable  $\tau$  is called an  $(\mathcal{F}_t)_{t\in T}$ -stopping time or  $(\mathcal{F}_t)_{t\in T}$ -optional time if  $\{\tau \leq t\} \in \mathcal{F}_t$  for all  $t \in T$ .  $A \ T \cup \{\infty\}$ -valued random variable  $\tau$  is called a weakly  $(\mathcal{F}_t)_{t\in T}$ -optional time if  $\{T < t\} \in \mathcal{F}_t$  for all  $t \in T$ .

We can collect several results:

- (1) Let  $(\mathcal{F}_t)_{t\in T}$  be a right continuous filtration, then every weak optional time is optional.
- (2) Suprema of sequences of optional times are optional, infima of sequences of weakly optional times are weakly optional.
- (3) Any weakly optional time  $\tau$  taking values in  $\mathbb{R} \cup \{\pm \infty\}$  can be approximated by some countably valued optional time  $\tau_n \searrow \tau$ , take for instance  $\tau_n := 2^{-n} [2^n \tau + 1]$ .

A stochastic process is a family of random variables  $(X_t)_{t\in T}$  on a filtered probability space  $(\Omega, \mathcal{F}, P)$ . The process is said to be  $(\mathcal{F}_t)_{t\in T}$ -adapted if  $X_t$  is  $\mathcal{F}_t$ measurable for all  $t \in T$ . Two stochastic processes  $(X_t)_{t\in T}$  and  $(Y_t)_{t\in T}$  are called versions of each other if  $X_t = Y_t$  almost surely for all  $t \in T$ , i.e. for every  $t \in T$ there is a null set  $N_t$  such that  $X_t(\omega) = Y_t(\omega)$  for  $\omega \notin N_t$ . Two stochastic processes  $(X_t)_{t\geq 0}$  and  $(Y_t)_{t\geq 0}$  are called *indistinguishable* if almost everywhere  $X_t = Y_t$  for all  $t \in T$ , i.e. there is a set N such that for all  $t \in T$  equality  $X_t(\omega) = Y_t(\omega)$  holds true for  $\omega \notin N$ . A stochastic process is called cadlag or RCLL (caglad or LCRL) if the sample paths  $t \mapsto X_t(\omega)$  are right continuous and have limits at the left side (left continuous and have limits at the right hand side). Given a stochastic process  $(X_t)_{t\in T}$  we denote by  $\hat{X}$  the associated mapping from  $T \times \Omega$  to  $\mathbb{R}$ , which maps  $(t, \omega)$  to  $X_t(\omega)$ .

Given a general filtration: a stochastic process  $(X_t)_{t\in T}$  is called progressively measureable or progressive if  $\hat{X} : ((T \cap [-\infty, t]) \times \Omega, \mathcal{B}(T \cap [-\infty, t]) \otimes \mathcal{F}_t) \to \mathbb{R}$  is measureable for all t. An adapted stochastic process with right continuous paths (or left continuous paths) almost surely is progressively measureable. A stochastic process  $(X_t)_{t\geq 0}$  is called measureable if  $\hat{X}$  is measureable with respect to the product  $\sigma$  algebra. A measureable adapted process has a progressively measureable modification, which is a complicated result.

Given a general filtration we can consider an associated filtration on the convex hull  $\overline{T}$  of T, where right from right-discrete points (i.e.  $t \in T$  such that there is  $\delta > 0$  with  $]t, t + \delta[\cap T = \emptyset)$ , a right continuous extension is performed. We can introduce two  $\sigma$ -algebras on  $\overline{T} \times \Omega$ , namely the one generated by left continuous processes (the predictable  $\sigma$ -algebra) and the one generated by the right continuous processes (the optional  $\sigma$ -algebra). Given a stochastic processes  $(X_t)_{t\in T}$  then we call it predictable if there is a predictable extension on  $\overline{T}$ .

Most important stopping times are hitting times of stochastic processes  $(X_t)_{t\in T}$ : given a Borel set B we can define the hitting time  $\tau$  of a stochastic process  $(X_t)_{t\in T}$  via

$$\tau = \inf\{t \in T \mid X_t \in B\}.$$

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There are several profound results around hitting times useful in potential theory, we provide some simple ones:

- (1) Let  $(X_t)_{t\in T}$  be an  $(\mathcal{F}_t)_{t\in T}$ -adapted right continuous process with respect to a general filtration and let B be an open set, then the hitting time is a weak  $(\mathcal{F}_t)_{t\in T}$ -stopping time.
- (2) Let  $(X_t)_{t\in T}$  be an  $(\mathcal{F}_t)_{t\in T}$ -adapted continuous process with respect to a general filtration and let B be a closed set, then the hitting time is a  $(\mathcal{F}_t)_{t\in T}$ -stopping time.
- (3) It is a very deep result in potential theory that all hitting times of Borel subsets are  $(\mathcal{F}_t)_{t\in T}$ -stopping times, if usual conditions are satisfied and the process is progressively measurable (Debut theorem).

The  $\sigma$ -algebra  $\mathcal{F}_t$  represents the set of all (theoretically) observable events up to time t including t, the stopping time  $\sigma$ -algebra  $\mathcal{F}_{\tau}$  represents all (theoretically) observable events up to  $\tau$ :

$$\mathcal{F}_{\tau} = \{ A \in \mathcal{F}, A \cap \{ \tau \le t \} \in \mathcal{F}_t \text{ for all } t \in T \}.$$

**Definition 2.2.** A stochastic process  $(X_t)_{t \in T}$  is called martingale (submartingale, supermartingale) with respect to a filtration  $(\mathcal{F}_t)_{t \in T}$  if

$$E[X_t \mid \mathcal{F}_s] = X_s$$

for  $t \geq s$  in T ( $E[X_t | \mathcal{F}_s] \geq X_s$  for submartingales,  $E[X_t | \mathcal{F}_s] \leq X_s$  for supermartingales).

**Proposition 2.3.** Let M be a martingale on an arbitrary index set T with respect to a filtration  $(\mathcal{G}_t)_{t\in T}$ . Assume a second filtration  $(\mathcal{F}_t)_{t\in T}$  such that  $\mathcal{F}_t \subset \mathcal{G}_t$  for  $t \in T$  and assume that M is actually also  $\mathcal{F}$  adapted, then M is also a martingale with respect to the filtration  $(\mathcal{F}_t)_{t\in T}$ .

*Proof.* The proof is a simple consequence of the tower property

$$\mathbb{E}[M_t | \mathcal{F}_s] = \mathbb{E}[\mathbb{E}[M_t | \mathcal{G}_s] | \mathcal{F}_s] = \mathbb{E}[M_s | \mathcal{F}_s] = M_s$$

for  $s \leq t$  in T.

Remember the Doob-Meyer decomposition for any integrable adapted stochastic process  $(X_t)_{t\in\mathbb{N}}$  on a *countable index set* T: there is a unique martingale M and a unique predictable process A with  $A_0 = 0$  such that X = M + A. It can be defined directly via

$$A_t := \sum_{s < t} E[X_{s+1} - X_s \mid \mathcal{F}_s]$$

for  $t \in \mathbb{N}$ , and M = X - A. The decomposition is unique since a predictable martingale starting at 0 vanishes. If X is a super-martingale, then the Doob-Meyer decomposition X = M - A yields a martingale and an decreasing process -A. We shall need in the sequel a curious inequality for bounded, non-negative supermartingales, see [4] (from where we borrow the proof). We state first a lemma for super-martingales of the type  $X_s = E[A_{\infty}|\mathcal{F}_s] - A_s$ , for  $s \ge 0$ , where A is a non-negative increasing process with limit at infinity  $A_{\infty}$  (such a super-martingale is called a potential). Assume  $0 \le X \le c$ , then

$$E[A^p_{\infty}] \le p! c^{p-1} E[X_0],$$

for natural numbers  $p \geq 1$ . Indeed

$$A_{\infty}^{p} = \sum_{i_{1},\dots,i_{p}} (A_{i_{1}+1} - A_{i_{1}}) \dots (A_{i_{p}+1} - A_{i_{p}})$$
  
=  $p \sum_{j} \sum_{j_{1},\dots,j_{p-1} \leq j} (A_{j_{1}+1} - A_{j_{1}}) \dots (A_{j_{p-1}+1} - A_{j_{p-1}})(A_{j_{1}+1} - A_{j})$   
=  $p \sum_{j_{1},\dots,j_{p-1}} (A_{j_{1}+1} - A_{j_{1}}) \dots (A_{j_{p-1}+1} - A_{j_{p-1}})(A_{\infty} - A_{j_{1}} \vee \dots \vee j_{p-1})$ 

Taking expectations, observing that all terms except the last one are measurable with respect to  $\mathcal{F}_{j_1 \vee \ldots \vee j_{p-1}}$  and inserting

$$X_{j_1 \vee \ldots \vee j_{p-1}} = E[A_{\infty} - A_{j_1 \vee \ldots \vee j_{p-1}} \mid \mathcal{F}_{j_1 \vee \ldots \vee j_{p-1}}] \le c$$

we obtain the recursive inequality

$$E[A^p_{\infty}] \le p E[A^{p-1}_{\infty}]c,$$

which leads by induction to the desired result. Taking now any non-negative supermartingale  $0 \le X \le c$ , then

$$(2.1) E[A_{\infty}^p] \le p! c^{p-1} E[X_0]$$

for natural numbers  $p \ge 1$ . Notice that for bounded super-martingales the limits to infinity exist due to martingale convergence results, see below Lemma 4.1.

Let us now consider a non-negative super-martingale X bounded by a constant  $c \ge 0$ , then we can apply the previous inequality for the potential  $X_0, \ldots, X_n, X_\infty, 0, 0, \ldots$  with corresponding  $\tilde{A}_{\infty} = A_n + E[X_n - X_{\infty} | \mathcal{F}_n] + X_{\infty} \ge A_n + X_{\infty}$ . If we let n tend to  $\infty$ , we obtain the result, namely that

$$E[M^p_\infty] \le p! c^{p-1} E[X_0] \,,$$

since  $M_{\infty} = A_{\infty} + X_{\infty}$ .

### 3. Optional Sampling: the discrete case

The most important theorem of this section is Doob's optional sampling theorem: it states that the stochastic integral, often also called martingale transform, with respect to a martingale is again a martingale. Most of the proofs stem from Olav Kallenberg's book [3].

Let us consider a finite index set T, a stochastic process  $X = (X_t)_{t \in T}$  and an increasing sequence of stopping times  $(\tau_k)_{k \geq 0}$  taking values in T together with bounded random variables  $V_k$ , which are  $\mathcal{F}_{\tau_k}$ , then

$$V_t := \sum_{k \ge 0} V_k \mathbf{1}_{\{\tau_k < t \le \tau_{k+1}\}}$$

is a predictable process. We call such processes *simple predictable* and we can define the stochastic integral (as a finite sum)

$$(V \bullet X)_{s,t} := \sum_{k \ge 0} V_k (X_{s \lor t \land \tau_{k+1}} - X_{s \lor t \land \tau_k}) \,.$$

By simple conditioning arguments we can prove the following proposition:

**Proposition 3.1.** Assume that the stopping times are deterministic and let X be a martingale, then  $(V \bullet X)$  is a martingale. If X is a sub-martingale and  $V \ge 0$ , then  $(V \bullet X)$  is a sub-martingale, too. Furthermore it always holds that  $(V \bullet (W \bullet X)) = (VW \bullet X)$  (Associativity).

This basic proposition can be immediately generalized by considering stopping times instead of deterministic times, since for any two stopping times  $\tau_1 \leq \tau_2$  taking values in T we have

$$1_{\{\tau_1 < t \le \tau_2\}} = \sum_{k \ge 0} 1_{\{\tau_1 < t \le \tau_2\}} 1_{\{t_k < t \le t_{k+1}\}}$$

for the sequence  $t_0 < t_1 < t_2 < ...$  exhausting the finite set T. Whence we can argue by Associativity  $(V \bullet (W \bullet X)) = (VW \bullet X)$  that we do always preserve the martingale property, or the sub-martingale property, respectively, in case of non-negative integrands.

We obtain several simple conclusions from these basic facts:

- (1) A stochastic process X is a martingale if and only if for each pair  $\sigma$ ,  $\tau$  of stopping times taking values in T we have that  $E[M_{\tau}] = E[M_{\sigma}]$ .
- (2) For any stopping time  $\tau$  taking values in  $T \cup \{\infty\}$  and any process M the stopped process  $M^{\tau}$

$$M_t^{\tau} := M_{\tau \wedge t}$$

for  $t \in T$  is well defined. If M is a (sub)martingale,  $M^{\tau}$  is a (sub)martingale, too, with respect to  $(\mathcal{F}_t)_{t\in T}$  and with respect to the stopped filtration  $(\mathcal{F}_{t\wedge\tau})_{t\in T}$ .

(3) For two stopping times  $\sigma$ ,  $\tau$  taking values in T, and for any martingale M we obtain

$$\mathbb{E}\left[\left.M_{\tau}\right|\mathcal{F}_{\sigma}\right] = M_{\sigma\wedge\tau}\,.$$

For this property we just use that  $(M_t - M_{t\wedge\tau})_{t\in T}$  is a martingale with respect to  $(\mathcal{F}_t)_{t\in T}$  as difference of two martingales. Stopping this martingale with  $\sigma$  leads to a martingale with respect to  $(\mathcal{F}_{t\wedge\sigma})_{t\in T}$ , whose evaluation at sup T and conditional expectation on  $\mathcal{F}_{\sigma}$  leads to the desired property.

These basic considerations are already sufficient to derive fundamental inequalities, from which the important set of maximal inequalities follows.

**Theorem 3.2.** Let X be a submartingale on a finite index set T with maximum  $supT \in T$ , then for any  $r \ge 0$  we have

$$rP[\sup_{t \in T} X_t \ge r] \le E[X_{\sup T} 1_{\{\sup_{t \in T} X_t \ge r\}}] \le E[X^+_{\sup T}]$$

and

$$rP[\sup_{t\in T} |X_t| \ge r] \le 3\sup_{t\in T} E[X_t].$$

*Proof.* Consider the stopping time  $\tau = \inf\{s \in T \mid X_s \ge r\}$  (which can take the value  $\infty$ ) and the predictable strategy

$$V_t := 1_{\{\tau \le \sup T\}} 1_{\{\tau < t \le \sup T\}},$$

then  $E[(V \bullet X)] \ge 0$  by Proposition 3.1, hence the first assertion follows. For the second one take X = M + A the Doob-Meyer decomposition of X and consider the first inequality for -M, hence

$$rP[\inf_{t \in T} X_t \le -r] \le E[M_{\sup T}^-] \le E[X_{\sup T}] - E[M_{\sup T}] \le 2\sup_{t \in T} E[X_t],$$

which combined with the first inequality yields the result.

**Theorem 3.3.** Let M be a martingale on a finite index set T with maximum  $\sup T$ 

$$E[\sup_{t \in T} |M_t|^p] \le \left(\frac{p}{p-1}\right)^p E[|M_{\sup T}|^p]$$

for all p > 1.

*Proof.* We apply the Levy-Bernstein inequalities from Theorem 3.2 to the submartingale |M|. This yields by Fubini's theorem and the Hölder inequality

$$\begin{split} E[\sup_{t\in T} |M_t|^p] &= p \int_0^\infty P[\sup_{t\in T} |M_t| > r] r^{p-1} dr \\ &\leq p \int_0^\infty E[|M_{\sup T}| \mathbf{1}_{\{\sup_{t\in T} |M_t| \ge r\}}] r^{p-2} dr \\ &= p E[|M_{\sup T}| \int_0^{\sup_{t\in T} |M_t|} r^{p-2} dr] \\ &= (\frac{p}{p-1}) E[|M_{\sup T}| \sup_{t\in T} |M_t|^{p-1}] \leq \|M_{\sup T}\|_p \|\sup_{t\in T} |M_t|^{p-1}\|_q \,, \end{split}$$
 ich yields the result.

which yields the result.

Finally we prove Doob's upcrossing inequality, which is the heart of all further convergence theorems, upwards or downwards. Consider an interval [a, b] and consider a submartingale X, then we denote by N([a, b], X) the number of upcrossings of a trajectory of X from below a to above b. We have the following fundamental lemma:

**Lemma 3.4.** Let X be s submartingale, then

$$E[N([a,b],X)] \le \frac{(X_{\sup T} - a)^+}{b-a}$$

*Proof.* Consider the submartingale  $(X - a)^+$  instead of X, since the number of upcrossings over [0, b-a] for  $(X-a)^+$  coincides with the number of upcrossings over [a, b] for X. Denote furthermore  $\tau_0 = \min T$ , then define recursively the stopping times

$$\sigma_k := \inf\{t \ge \tau_{k-1} \mid X_t = 0\}$$

and

$$\tau_k := \inf\{t \ge \sigma_k \mid X_t \ge b - a\}$$

for  $k \geq 1$ . The process

$$V_t := \sum_{k \ge 1} \mathbf{1}_{\{\tau_k < t \le \sigma_k\}}$$

is predictable and  $V \ge 0$  such as 1 - V. We conclude by

$$(b-a)E[N([a,b],(X-a)^+)] \le E[(V \bullet (X-a)^+)_{\sup T}] \le E[(1 \bullet (X-a)^+)_{\sup T}] \le E[(X_{\sup T}-a)^+].$$

Mind the slight difference of these inequalities to the one in the introduction. 

This remarkable lemma allows to prove the following deep convergence results by passing to countable index sets:

**Theorem 3.5.** Let X be an  $L^1$  bounded submartingale on a countable index set T, then there is a set A with probability one such that  $X_t$  converges along any increasing or decreasing sequence in T.

*Proof.* By  $\sup_{t \in T} E[|X_t|] < \infty$  we conclude by the Lévy-Bernstein inequality that the measurable random variable  $\sup_{t \in T} |X_t|$  is finitely valued, hence along every subsequence there is a finite inferior or superior limit. By monotone convergence we know that for any interval the number of upcrossings is finite almost surely. Consider now A, the intersection of sets with probability one, where the number of upcrossings is finite over intervals with rational endpoints. A has again probability one and on A the process X converges along any increasing or decreasing subsequence, since along monotone sequences a finite number of upcrossings leads to equal inferior and superior limits. Notice that we work here with monotone convergence, since the number of upcrossings for increasing index sets is increasing, however, its expectation is bounded.

**Theorem 3.6.** For any martingale M on any index set we have the following equivalence:

- (1) M is uniformly integrable.
- (2) M is closeable at sup T.
- (3) M is  $L^1$  convergent at sup T.

*Proof.* If M is closeable on an arbitrary index set T, then by definition there is  $\xi \in L^1(\Omega)$  such that  $M_t = E[\xi | \mathcal{F}_t]$  for  $t \in T$ , hence

$$E[M_t 1_A] \le E[E[|\xi| | \mathcal{F}_t] 1_A] = E[|\xi| E[1_A | \mathcal{F}_t]]$$

for any  $A \in \mathcal{F}$ , which tends to zero if  $P(A) \to 0$ , uniformly in t, hence uniform integrability. On the other hand a uniformly integrable martingale is bounded in  $L^1$  and therefore we have one and the same almost sure limit along any subsequence increasing to sup T. If M is uniformly integrable, an almost sure limit is in fact  $L^1$ .

Finally assume  $M_t \to \xi$  for  $t \to \sup T$  in  $L^1$ , hence  $M_s \to E[\xi | \mathcal{F}_s]$ , for  $s \in T$ and the martingale property, hence  $M_s = E[\xi | \mathcal{F}_s]$  for any  $s \in T$ , which concludes the proof.

We obtain the following beautiful corollary:

**Corollary 3.7.** Let M be a martingale on an arbitrary index set and assume p > 1, then  $M_t$  converges in  $L^p$  for  $t \to \sup T$  if and only if it is  $L^p$  bounded.

*Proof.* If M is  $L^p$  bounded, then it is uniformly integrable (by Doob's maximal inequalities from Theorem 3.3) and convergence takes place in  $L^1$  by the previous theorem, which in turn by  $L^p$ -boundedness is also a convergence in  $L^p$ . On the other hand, if M converges in  $L^p$ , then it is by the Jensen's inequality also  $L^p$  bounded.

Finally we may conclude the following two sided version of closedness:

**Theorem 3.8.** Let T be a countable index set unbounded above and below, then for any  $\xi \in L^1$  we have that

$$E[\xi \mid \mathcal{F}_t] \to E[\xi \mid \mathcal{F}_{\pm \infty}]$$

for  $t \to \pm \infty$ .

*Proof.* By  $L^1$  boundedness we obtain convergence along any increasing or decreasing subsequence towards limits  $M_{\pm\infty}$ . The upwards version follows from the previous Theorem 3.6, the downwards version follows immediately.

This last theorem is often called reserve martingale convergence and can be used to prove the strong law of large numbers and almost sure convergence of quadratic variations:

(1) Let  $(\xi_i)_{i\geq 1}$  be an i.i.d. sequence of random variables in  $L^1$ . Then we can consider the filtration

$$\mathcal{F}_{-n} := \sigma(S_n, S_{n+1}, \ldots)$$

with  $S_n := \sum_{i=1}^n \xi_i$ , for  $n \ge 1$ . Since  $(\xi_1, S_n, S_{n+1}, \ldots) = (\xi_k, S_n, S_{n+1}, \ldots)$  for  $k \le n$  in distribution, we obtain

$$E[\xi_1 \mid \mathcal{F}_{-n}] = E[\xi_k \mid \mathcal{F}_{-n}]$$

for  $k \leq n$ . Whence

$$\frac{S_n}{n} = \frac{1}{n} E[S_n \mid \mathcal{F}_{-n}] = \frac{1}{n} \sum_{i=1}^n E[\xi_i \mid \mathcal{F}_{-n}]$$
$$= \frac{1}{n} \sum_{i=1}^n E[\xi_1 \mid \mathcal{F}_{-n}] = E[\xi_1 \mid \mathcal{F}_{-n}] \to E[\xi_1].$$

since the intersection of all tail  $\sigma$  algebras is trivial, i.e. all elements of the intersection have probability either 0 or 1 (Hewitt-Savage 0 - 1 law).

(2) Let  $(W_t)_{t \in [0,1]}$  be a Wiener process and denote – for a fixed time  $0 \le t \le 1$ – by

$$V_t^n := \sum_{t_i \in \Pi^n} (W_{t_{i+1}} - W_{t_i})^2$$

the approximations of quadratic variation t along a refining sequence of partitions  $\Pi^n \subset \Pi^{n+1}$  of [0, t], whose meshes tend to zero. We know from the lecture notes that the approximation takes place in  $L^2$ , but we do not know whether it actually holds almost surely. Consider the filtration

$$\mathcal{F}_{-n} := \sigma(V_t^n, V_t^{n+1}, \ldots)$$

for  $n \geq 1$ , whose intersection is actually trivial. Without loss of generality we assume that each partition  $\Pi^n$  contains n partition points by possibly adding to the sequence intermediate partitions. Fix now  $n \geq 2$ , then we consider the difference between  $\Pi^{n-1}$  and  $\Pi^n$ , which is just one point vlying only in  $\Pi^n$  and being surrounded by two nearest neighboring points from  $\Pi^{n-1}$ , i.e. u < v < w. Consider now a second Wiener process  $\tilde{W}_s =$  $W_{s\wedge v} - (W_s - W_{s\wedge v})$ . Apparently

$$(\tilde{V}_t^{n-1}, V_t^n, V_t^{n+1}, \dots) = (V_t^{n-1}, V_t^n, V_t^{n+1}, \dots)$$

in distribution, hence it holds that

$$E[\tilde{V}_t^{n-1} - V_t^n \mid \mathcal{F}_{-n}] = E[V_t^{n-1} - V_t^n \mid \mathcal{F}_{-n}],$$

which in turn means that

$$E[(\tilde{W}_u - \tilde{W}_v)(\tilde{W}_v - \tilde{W}_w) \mid \mathcal{F}_{-n}] = E[(W_u - W_v)(W_v - W_w) \mid \mathcal{F}_{-n}].$$

Inserting the definition of  $\tilde{W}$  yields the result that

$$E[V_t^{n-1} - V_t^n \mid \mathcal{F}_{-n}] = 0$$

for  $n \geq 1$ . Hence we have a martingale on the index set  $\mathbb{Z}_{\leq 1}$ , which by martingale convergence tends almost surely to its  $L^2$  limit t.

# 4. Martingales on continuous index sets

Martingale inequalities on uncountable index sets can often be derived from inequalities for the case of countable index sets if certain path properties are guaranteed. From martingale convergence results on countable index sets we can conclude the existence of RCLL versions for processes like martingales, which is the main result of this section. Most of the proofs stem from Olav Kallenberg's book [3].

We need an auxiliary lemma on reverse submartingales first. Of course similar statements hold for supermartingales.

**Lemma 4.1.** Let X be a submartingale on  $\mathbb{Z}_{\leq 0}$ . Then X is uniformly integrable if and only if E[X] is a bounded (from below) sequence.

*Proof.* Let E[X] be bounded from below. We can then introduce a Doob-Meyer type decomposition, i.e.

$$A_n := \sum_{k < n} E[X_{k+1} - X_k \mid \mathcal{F}_k],$$

which is well defined since all summands are positive due to submartingality and

$$E[A_0] \le E[X_0] - \inf_{n \ge 0} E[X_n] < \infty.$$

Whence X = M + A, where M is a martingale. Since A is uniformly integrable and M is a martingale being closed at 0 by martingale convergence, hence uniformly integrable, also the sum is uniformly integrable. The other direction follows immediately since  $E[X_n]$  is decreasing for  $n \to \infty$ . If it were unbounded from below, it cannot be uniformly integrable.

From this statement we can conclude by martingale convergence the following fundamental regularization result:

**Theorem 4.2.** For any submartingale X on  $\mathbb{R}_{\geq 0}$  with restriction Y to  $\mathbb{Q}_{\geq 0}$  we have:

- (1) The process of right hand limits  $Y^+$  exists on  $\mathbb{R}_{\geq 0}$  outside some nullset A and  $Z := \underline{1}_{A^c}Y^+$  is an RCLL submartingale with respect to the augmented filtration  $\overline{\mathcal{F}_+}$ .
- (2) If the filtration is right continuous, then X has a RCLL version, if and only if  $t \mapsto E[X_t]$  is right continuous.

*Proof.* The process Y is  $L^1$  bounded on bounded intervals since the positive part is an  $L^1$  bounded submartingale by Jensen's inequality, hence by martingale convergence Theorem 3.5 we obtain the existence of right and left hand limits and therefore  $Y^+$  is RCLL. Clearly the process Z is adapted to the augmented filtration  $\overline{\mathcal{F}_+}$ .

The submartingale property follows readily, too: fix times s < t and choose  $s_n \searrow s$  and  $t_m \searrow t$ , with  $s_n < t$  for all  $n \ge 1$ . Then – by assumption –  $E[Y_{t_m} | \mathcal{F}_{s_n}] \ge Y_{s_n}$ . By martingale convergence to the left we obtain

$$E[Y_{t_m} \mid \mathcal{F}_{s+}] \ge Z_s$$

almost surely. Since the submartingale  $(Y_{t_m}))_{m\geq 1}$  has bounded expectations, we conclude  $L^1$ -convergence (due to uniform integrability by the previous lemma) and therefore arrive at

$$E[Z_t \mid \overline{\mathcal{F}_{s+}}] \ge Z_s$$

For the second assertion observe that if X is RCLL, then the curve E[X] is right continuous by uniform integrability along decreasing subsequences and the previous lemma. On the other hand if E[X] is right continuous  $Z_t = E[Z_t | \mathcal{F}_t] \ge X_t$  by limits from the right, but  $E[Z_t - X_t] = 0$  by right continuity of E[X], hence Z and X are indistinguishable.

### 5. Stochastic Integration for Caglad Integrands

We follow here mainly Philip Protter's book [5] on stochastic integration, which is inspired by works of Klaus Bichteler, Claude Dellacherie and Paul-Andre Meyer, see [2] and [4]: the idea is to crystallize the essential property (the "good integrator" property), which guarantees the existence of stochastic integrals, and to derive all properties of stochastic integrals from the good integrator property. Finally it can be shown that every good integrator is the sum of a local martingale and a finite variation process. This approach leads to an integration theory for caglad integrands.

Let us introduce some notation: we denote by  $\mathbb{S}$  the set of simple predictable processes, i.e.

$$H_0 1_{\{0\}} + \sum_{i=1}^n \sum_i H_i 1_{]T_i, T_{i+1}]}$$

for an increasing, finite sequence of stopping times  $0 = T_0 \leq T_1 \leq \ldots T_{n+1} < \infty$ and  $H_i$  being  $\mathcal{F}_{T_i}$  measurable, by  $\mathbb{L}$  the set of caglad processes and by  $\mathbb{D}$  the set of cadlag processes on  $\mathbb{R}_{>0}$ . These vector spaces are endowed with the metric

$$d(X,Y) := \sum_{n \ge 0} \frac{1}{2^n} E\big[ |(X-Y)|_n^* \wedge 1 \big],$$

which makes  $\mathbb{L}$  and  $\mathbb{D}$  complete topological vector spaces. We call this topology the ucp-topology ("uniform convergence on compacts in probability"). Notice that predictable strategies as well as integrators are considered  $\mathbb{R}$  valued here, which, however, *contains* the  $\mathbb{R}^n$  case.

Notice that we are dealing here with topological vector spaces, which are not even locally convex. This leads also to the phenomenon that the metric does not detect boundedness of sets, which is defined in the following way: A subset B of a topological vector space  $\mathbb{D}$  is called bounded, if it can be absorbed by any open neighborhood U of zero, i.e. there is R > 0 such that  $B \subset RU$ . For instance for the space of random variables  $L^0(\Omega)$  this translates to the following equivalent statement: a set B of random variables is bounded in probability if for every  $\epsilon > 0$ there is c > 0 such that

$$P[|Y| \ge c] < \epsilon$$

for  $Y \in B$ , which is of course not detectable by the metric.

In order to facilitate the reasoning we shall use the following definition, which seems slightly less general than the original ones which is treated in the subsequent remark. **Definition 5.1.** A cadlag process X is called good integrator if the map

$$J_X:\mathbb{S}\to\mathbb{D}$$

with  $J_X(H) := H_0 X_0 + \sum_{i=1}^n H_i (X_{T_{i+1} \wedge t} - X_{T_i \wedge t})$ , for  $H \in \mathbb{S}$ , is continuous with respect to the ucp-topologies.

Remark 5.2. It would already be sufficient to require a good integrator to satisfy the following property: for every stopped process  $X^t$  the map  $I_{X^t} : \mathbb{S}_u \to L^0(\Omega)$ , where  $I_{X^t}(H) := J_X(H)_{\infty}$ , from simple predictable processes with the uniform topology  $\mathbb{S}_u$  to random variables with convergence in probability, is continuous, i.e.

(5.1) 
$$I_{X^{t}}(H^{k}) := H_{0}^{k}X_{0} + \sum_{i=1}^{n} H_{i}^{k}(X_{T_{i+1}\wedge t} - X_{T_{i}\wedge t}) \to 0$$

if  $H^k \to 0$  uniformly on  $\Omega \times \mathbb{R}_{\geq 0}$ .

Clearly Property (5.1) holds if it only holds locally: indeed let  $\tau^n$  be a localizing sequence, i.e.  $\tau^n \nearrow \infty$  with  $X^{\tau^n}$  being a good integrator. Fix  $t \ge 0$  and a sequence  $H^k \to 0$ , which tends to 0 uniformly, then

$$P[|I_{X^{t}}(H^{k})| \ge c] \le P[|I_{X^{\tau^{n} \wedge t}}(H^{k})| \ge c] + P[\tau^{n} \le t]$$

for every  $n \ge 0$ . Hence we can choose *n* large enough, such that the second term is small by the localizing property, and obtain for *k* large enough that the first term is small by  $X^{\tau^n}$  being a good integrator.

From continuity with respect to the uniform topology continuity with respect to the ucp topologies, as claimed in the definition of a good integrator, immediately follows. Indeed assume that  $I_X$  is continuous with respect to the uniform topology, fix  $n \ge 0$  and a sequence  $H^k \to 0$ , which tends to 0 uniformly, then choose  $c \ge 0$  and define a sequence of stopping times

$$\tau^k := \inf\{t \mid |(H^k \bullet X)_t| \ge c\}$$

for  $k \geq 0$ , then

$$P[|(H^{k} \bullet X)_{n}|^{*} \ge c] \le P[|(H^{k} 1_{[0,\tau^{k}]} \bullet X)_{n}| \ge c] \to 0$$

as  $k \to \infty$  by assumption, hence  $J_X : \mathbb{S}_u \to \mathbb{D}$  is continuous. Take now a sequence  $H^k \to 0$  in ucp, and choose  $c \ge 0$ ,  $\epsilon > 0$  and  $n \ge 0$ . Then there is some  $\eta > 0$  such that

$$P[|(H \bullet X)_n|^* \ge c] \le \epsilon$$

for  $||H||_{\infty} \leq \eta$ . Define furthermore stopping times

$$\rho^k := \inf\{s \mid |H_s^k| > \eta\}$$

then we obtain

$$P[|(H^k \bullet X)_n|^* \ge c] \le P[|(H^k \mathbf{1}_{]0,\rho^k]} \bullet X)_n| \ge c] + P[\rho^k < n] < 2\epsilon$$

if k is large enough since  $P[\rho^k < n] \to 0$  as  $k \to \infty$ .

*Remark* 5.3. The set S is dense in the ucp-topology in  $\mathbb{L}$ , even the bounded simple predictable processes are dense. Consider just the sequence of partitions introduced at the beginning of the next subsection.

Remark 5.4. We provide examples of good integrators:

• Any process of finite variation A is a good integrator, since for every simple (not even predictable) process it holds that

$$\left|\int_{0}^{t} H_{s} dA_{s}\right| \leq \left\|H\right\|_{\infty} \int_{0}^{t} d|A|_{s}$$

almost surely, for  $t \ge 0$ .

• By Ito's fundamental insight square integrable martingales M are good integrators, since

$$E\left[(H \bullet M)_t^2\right] \le \left\|H\right\|_{\infty} E\left[\left|M_t\right|^2\right]$$

holds true for simple, bounded and predictable processes  $H \in b\mathbb{S}$ .

• By the following elementary inequality due to Burkholder we can conclude that martingales are good integrators: for every martingale M and every simple, bounded and predictable process  $H \in bS$  it holds that

$$cP(|(H \bullet M)|_{1}^{*} \ge c) \le 18||H||_{\infty}||M_{1}||_{1}$$

for all  $c \geq 0$ . For an easy proof of this inequality see, e.g., works of Klaus Bichteler [2] or Paul-Andre Meyer [4, Theorem 47, p. 50]. Since the inequality is crucial for our treatment, we shall prove it here, too. Notice that we are just dealing with integrals with respect to simple integrands, hence we can prove it for discrete martingales on a finite set of time points. Let M be a non-negative martingale first and H bounded predictable with  $\|H\|_{\infty} \leq 1$ , then  $Z := M \wedge c$  is a supermartingale and we have

$$cP(|(H \bullet M)|_1^* \ge c) \le cP(|M|_1^* \ge c) + cP(|(H \bullet Z)|_1^* \ge c)$$
.

Since Z is a super-martingale we obtain by the Doob-Meyer decomposition for discrete super-martingales  $Z = \tilde{M} - A$  that

$$(H \bullet Z) \le (H \bullet \tilde{M}) + A,$$

i.e. we have an upper bound being a sub-martingale. With  $(H \bullet \tilde{M}) + A$  also its square is a sub-martingale. Hence we can conclude by Lemma 3.2 that

$$cP(|(H \bullet M)|_1^* \ge c) \le E[M_1] + 2\frac{1}{c}E[(H \bullet \tilde{M})_1^2 + A_1^2],$$

since

$$cP(|(H \bullet Z)|_{1}^{*} \ge c) \le \le \le \frac{1}{c}P(|(H \bullet \tilde{M})_{1} + A_{1}|^{2} \ge c) \le 2\frac{1}{c}E[(H \bullet \tilde{M})_{1}^{2} + A_{1}^{2}]$$

Ito's insight allows to estimate the variance of the stochastic integral at time 1 by  $E[\tilde{M}_1^2]$ . Both quantities  $\tilde{M}$  and A of the Doob-Meyer decomposition may however be estimated through  $E[A_1^2] \leq E[\tilde{M}_1^2] \leq 2cE[Z_0] \leq 2cE[M_0]$ , see (2.1), since Z is non-negative (so  $A \leq M$  holds true) and  $Z \leq c$ . This leads to an upper bound

$$cP(|(H \bullet M)|_1^* \ge c) \le 9E[M_0].$$

Writing a martingale as difference of two non-negative martingales leads to the desired result. Apparently the result translates directly to the fact that M is a good integrator. We actually immediately obtain that  $J_X : \mathbb{S}_u \to \mathbb{D}$ 

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is continuous, wherefrom - as we have seen before - the continuity even with respect to the ucp topology on S follows.

By density and continuity we can extend the map  $J_X$  to all caglad processes  $Y \in \mathbb{L}$ , which defines the stochastic integral  $(Y \bullet X)$ . As a simple corollary we can prove the following proposition:

**Proposition 5.5.** Let  $H, G \in \mathbb{L}$  be given and let X be a good integrator, then  $(G \bullet X)$  is a good integrator and  $(H \bullet (G \bullet X)) = (HG \bullet X)$ .

5.1. Approximation results. Most important for the calculation and understanding of stochastic integrals is the following approximation result: a sequence of partition tending to identity  $\Pi^k$  consists of stopping times  $0 = T_0^k \leq \ldots \leq T_{i_k}^k < \infty$ with mesh  $\sup_i(T_{i+1}^k - T_i^k)$  tending to 0 and  $\sup_i T_i^k \to \infty$ . We call the sequence of cadlag processes

$$Y^{\Pi^k} := \sum_i Y_{T^k_i} 1_{]T^k_i, T^k_{i+1}]}$$

a sampling sequence for a cadlag process Y along  $\Pi^k$ , for  $k \ge 0$ .

*Example* 5.6. One important sequence of partition is the following one: let  $Y \in \mathbb{D}$  be a RCLL process. For  $n \ge 0$  we can define a double sequence of stopping times  $\tau_i^n$ 

$$\tau_0^n := 0 \text{ and } \tau_{i+1}^n := \inf\{s \ge \tau_i^n \mid |Y_s - Y_{\tau_i^n}| \ge \frac{1}{2^n}\}$$

for  $i \ge 0$ . This defines a sequence of partitions  $\Pi^n$  tending to identity. Apparently we have that

$$|Y_{-} - Y^{\Pi^n}| \le \frac{1}{2^n},$$

which means in particular that  $Y^{\Pi^n} \to Y_-$  in the ucp topology.

**Theorem 5.7.** For any good integrator X we obtain that

$$(Y_{-}^{\Pi^k} \bullet X) \to (Y_{-} \bullet X)$$

in the ucp topology in general, as well as the less usually stated but equally true ucp convergence result  $% \left( \frac{1}{2} \right) = 0$ 

$$(Y_{-}^{\Pi^{k}} \bullet X^{\Pi^{k}}) \to (Y_{-} \bullet X)$$

if  $X^{\Pi^n} \to X$  in the ucp topology.

*Proof.* We know by previous remarks that there are sequences  $Y^l$  of simple predictable processes converging ucp to Y, also  $Y_-^l \to Y_-$  holds true for the associated left continuous processes. Hence we can write

$$((Y_{-} - Y_{-}^{\Pi^{k}}) \bullet X) = ((Y_{-} - Y_{-}^{l}) \bullet X) + ((Y_{-}^{l} - (Y^{l})_{-}^{\Pi^{k}}) \bullet X) + (((Y^{l})_{-}^{\Pi^{k}} - Y_{-}^{\Pi^{k}}) \bullet X) + (((Y^{l})_{-}^{\Pi^{k}} - Y_{-}^{\Pi^{k}}) \bullet X) + (((Y^{l})_{-}^{\Pi^{k}}) \bullet X) + (((Y^{l})_{-}) \bullet X) + (((Y^{l})_{-})$$

where the first and third term converge of course in ucp, the third even uniformly in k. The middle term is seen to converge by direct inspection.

As a consequence we obtain of course for two good integrators X, Y that

$$[X^{\Pi^k}, Y^{\Pi^k}] \to [X, Y],$$

since quadratic co-variation can be expressed by stochastic integrals, i.e.

$$[X, X] := X^2 - 2(X_- \bullet X).$$

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Let us make a remark on important notations here: we shall always assume that  $X_{0-} = 0$  (a left limit coming from negative times), whereas  $X_0$  can be different from zero, whence  $\Delta X_0$  is not necessarily vanishing.

5.2. **Ito's theorem.** It is remarkable that Ito's theorem can be concluded from its version for piece-wise constant processes due to the following continuity lemma, which complements results which have already been established for the approximation of stochastic integrals. We state an additional continuity lemma:

**Lemma 5.8.** Let  $X^1, \ldots, X^n$  be good integrators,  $\Pi^k$  a sequence of partitions tending to the identity and  $f : \mathbb{R}^n \to \mathbb{R}$  a  $C^2$  function, then for  $t \ge 0$ 

$$\sum_{s \le t} \left\{ f(X_s^{\Pi^k}) - f(X_{s-}^{\Pi^k}) - \sum_{i=1}^n \partial_i f(X_{s-}^{\Pi^k}) \Delta X_s^{i,\Pi^k} - \frac{1}{2} \sum_{i,j=1}^n \partial_{ij}^2 f(X_{s-}^{\Pi^k}) \Delta X_s^{i,\Pi^k} \Delta X_s^{j,\Pi^k} \right\}$$
  
$$\to_{k \to \infty} \sum_{s \le t} \left\{ f(X_s) - f(X_{s-}) - \sum_{i=1}^n \partial_i f(X_{s-}) \Delta X_s^i - \frac{1}{2} \sum_{i,j=1}^n \partial_{ij}^2 f(X_{s-}) \Delta X_s^i \Delta X_s^j \right\},$$

where the limit can be even understood in ucp topology.

*Remark* 5.9. Here we mean with  $X^{\Pi^k}$  the RCLL version of the sampled process with a slight abuse of notation.

*Proof.* The proof of this elementary lemma relies on Taylor expansion of f: apparently the summands are small at s if  $\Delta X_s^{\Pi^k}$  is small, hence only those jumps remain after the limit, which are at time points where X actually jumps. Let us make this precise: first we know – by the very existence of quadratic variation – that

$$\sum_{s \leq t} \left( \Delta X^i \right)^2 \leq [X^i, X^i]_t < \infty$$

almost surely. Fix  $t \ge 0$  and  $\epsilon > 0$ , then we find for every  $\omega \in \Omega$  a finite set  $A_{\omega}$  of times up to t, where X jumps, and a possibly countable set of times  $B_{\omega}$  up to t, where X jumps and  $\sum_{s \in B} \|\Delta X\|^2 \le \epsilon^2$ , since every cadlag path has at most countably many jumps up to time t. Furthermore we know that

$$f(y) - f(x) - \sum_{i=1}^{n} \partial_i f(x)(y-x)^i - \frac{1}{2} \sum_{i,j=1}^{n} \partial_{ij}^2 f(x)(y-x)^i (y-x)^j = o(\|y-x\|) \|y-x\|^2$$

as  $y \to x$ . This means that for  $\omega \in \Omega$  we can split the sum into two sums denoted by  $\sum_A$ , where  $]T_i^k(\omega), T_{i+1}^k(\omega)] \cap A_\omega \neq \emptyset$ , and  $\sum_B$  else. We then obtain an estimate for the sum  $\sum_B$  of the type

$$\sum_{B} \leq 2\epsilon^2 o\bigl(\epsilon\bigr)$$

for k large enough by uniform continuity of continuous functions on compact intervals and jump size at most  $\epsilon$ . The other part  $\sum_A$  behaves differently, but is a finite sum, hence it converges to the respective limit (written with A in respective sense). Letting now tend  $\epsilon \to 0$  the result follows immediately. The argument is true uniformly along paths in probability.

We are now able to prove Ito's formula in all generality:

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**Theorem 5.10.** Let  $X^1, \ldots, X^n$  be good integrators and  $f : \mathbb{R}^n \to \mathbb{R}$  a  $C^2$  function, then for  $t \ge 0$ 

$$f(X_t) = f(X_0) + \sum_{i=1}^n (\partial_i f(X_-) \bullet X^i)_t + \frac{1}{2} \sum_{i,j=1}^n (\partial_{ij}^2 f(X_-) \bullet [X^i, X^j])_t + \sum_{0 < s \le t} \left\{ f(X_s) - f(X_{s-}) - \sum_{i=1}^n \partial_i f(X_{s-}) \Delta X_s^i - \frac{1}{2} \sum_{i,j=1}^n \partial_{ij}^2 f(X_{s-}) \Delta X_s^i \Delta X_s^j \right\}.$$

*Proof.* Let  $\Pi^k$  be a sequence of partitions tending to the identity, such that  $X^{\Pi^k} \to X$  in the ucp topology, then Ito's formula reads by careful inspection

$$\begin{split} f(X_t^{\Pi^k}) &= f(X_0) + \sum_{i=1}^n \left(\partial_i f(X_-^{\Pi^k}) \bullet X^{i,\Pi^k}\right)_t + \frac{1}{2} \sum_{i,j=1}^n \left(\partial_{ij}^2 f(X_-^{\Pi^k}) \bullet [X^{i,\Pi^k}, X^{j,\Pi^k}]\right)_t + \\ &+ \sum_{0 < s \le t} \left\{ f(X_s^{\Pi^k}) - f(X_{s-}^{\Pi^k}) - \sum_{i=1}^n \partial_i f(X_{s-}^{\Pi^k}) \Delta X_s^{i,\Pi^k} - \frac{1}{2} \sum_{i,j=1}^n \partial_{ij}^2 f(X_{s-}^{\Pi^k}) \Delta X_s^{i,\Pi^k} \Delta X_s^{j,\Pi^k} \right\} \end{split}$$

since the process is piece-wise constant and the sum is just telescoping. By the previously stated convergence result, however, this translates directly – even in ucp convergence – to the limit for  $k \to \infty$ , which is Ito's formula.

5.3. Quadratic Pure jump good integrators. We call a good integrator quadratic pure jump if  $[X, X]_t = \sum_{s \leq t} (\Delta X_s)^2$  for  $t \geq 0$ . It follows from Ito's formula that every cadlag, adapted and finite variation process X is quadratic pure jump.

Indeed the finite variation property yields a well-known Ito formula for  $f(x) = x^2$ (notice that the second order terms is missing in the sum) of the type

$$X_t^2 = X_0^2 + 2(X_- \bullet X) + \sum_{0 < s \le t} \left\{ X_s^2 - X_{s-}^2 - 2X_{s-} \Delta X_s \right\} = 2(X_- \bullet X) + \sum_{s \le t} \left( \Delta X_s \right)^2 + \sum_{s \ge t} \left( \Delta X_s \right)^2 + \sum_{s \le t} \left( \Delta X_s \right)^2 + \sum_{s \ge t} \left( \Delta X_s \right)^2$$

which yields the result on the quadratic variation. Hence for every good integrator M we obtain

$$\left[X,M\right]_t = \sum_{s \leq t} \Delta X \Delta M$$

for finite variation processes X with complete analogous arguments.

5.4. Stochastic exponentials. An instructive example how to calculate with jump processes is given by the following process: let X be a good integrator with  $X_0 = 0$ , then the process

$$Z_{t} = \exp\left(X_{t} - \frac{1}{2}[X, X]_{t}\right) \prod_{0 \le s \le t} (1 + \Delta X_{s}) \exp\left(-\Delta X_{s} + \frac{1}{2}(\Delta X_{s})^{2}\right)$$

satisfies  $Z_t = 1 + (Z_- \bullet X)_t$  and is called stochastic exponential.

For the proof we have to check that the infinite product is actually converging and defining a good integrator. We show this by proving that it defines an adapted, cadlag process of finite variation. We only have to check this for jumps smaller than  $\frac{1}{2}$ , i.e. we have to check whether

$$\sum_{s \le t} \{ \log(1 + U_s) - U_s + \frac{1}{2} U_s^2 \}$$

converges absolutely, where  $U_s := \Delta X_s \mathbf{1}_{\{|\Delta X_s| \ge \frac{1}{2}\}}$ , for  $s \ge 0$ . This, however, is true since  $|\log(1+x) - x + \frac{1}{2}| \le Cx^3$  for  $|x| \le \frac{1}{2}$  and  $\sum_{s \le t} \Delta X_s^2 \le [X, X] < \infty$  almost surely. Hence we can apply Ito's formula for the function  $\exp(x_1)x_2$  with good integrators

$$X_t^1 = X_t - \frac{1}{2} [X, X]_t$$

and

$$X_t^2 = \prod_{0 \le s \le t} (1 + \Delta X_s) \exp\left(-\Delta X_s + \frac{1}{2} (\Delta X_s)^2\right).$$

This leads to

$$Z_{t} = 1 + (Z_{-} \bullet X)_{t} - \frac{1}{2}(Z_{-} \bullet [X, X])_{t} + (\exp(X_{-}^{1}) \bullet X^{2})_{t} + \frac{1}{2}(Z_{-} \bullet [X^{1}, X^{1}])_{t} + \sum_{s \leq t} \{Z_{s} - Z_{s-} - Z_{s-} \Delta X_{s}^{1} - \exp(X_{s-}^{1}) \Delta X_{s}^{2} - \frac{1}{2}Z_{s-} (\Delta X^{1})^{2}\} = 1 + (Z_{-} \bullet X)_{t},$$

since

$$\Delta Z_s = Z_{s-} \exp\left(\Delta X_s - \frac{1}{2} (\Delta X_s)^2\right) (1 + \Delta X_s) \exp\left(-\Delta X_s + \frac{1}{2} (\Delta X_s)^2\right)$$

holds true, for  $s \ge 0$ .

5.5. Lévy' Theorem. Another remarkable application is Lévy's theorem: consider local martingales  $B^1, \ldots, B^n$  starting at 0 with continuous trajectories such that  $[B^i, B^j]_t = \delta^{ij} t$  for  $t \ge 0$ . Then  $B^1, \ldots, B^n$  are standard Brownian motions.

For this theorem we have to check that stochastic integrals along locally square integrable martingales are locally square integrable martingales: indeed let X be a locally square-integrable martingale and  $H \in \mathbb{L}$ , then by localizing and the formula

$$(H \bullet X)^{\tau} = (H \bullet X^{\tau}) = (H1_{]0,\tau]} \bullet X^{\tau}) = (H1_{]0,\tau]} \bullet X)$$

we can assume that H is bounded and X is in fact a square integrable martingale with  $E[X_{\infty}^2] < \infty$ . Then, however, we have by Ito's insight that for sequence of partitions tending to identity the process

$$(H^{\Pi^{\kappa}} \bullet X)$$

is a square integrable martingale, which satisfies additionally

$$E[(H^{\Pi^k} \bullet X)_{\infty}^2] \le \|H\|_{\infty} E[X_{\infty}^2].$$

By martingale convergence this means that the limit in probability of the stochastic integrals is also a square integrable martingale, whence the result.

This can be readily applied to the stochastic process

$$M_t := \exp\left(i\langle\lambda, B_t\rangle + \frac{t}{2}\|\lambda\|^2\right)$$

for  $t \ge 0$  and  $\lambda \in \mathbb{R}^n$ , which is a bounded local martingale by the previous consideration and Ito's formula. A bounded local martingale is a martingale, whence

$$E[\exp(\mathsf{i}\langle\lambda, B_t - B_s\rangle) | \mathcal{F}_s] = \exp(-\frac{t-s}{2} \|\lambda\|^2).$$

## 6. STOCHASTIC INTEGRATION FOR PREDICTABLE INTEGRANDS

For many purposes (i.e. martingale representation) it is not enough to consider only caglad integrands, but predictable integrands are needed. This cannot be achieved universally for all good integrators, but has to be done case by case. The main tool for this purpose are  $\mathcal{H}^p$  spaces, for  $1 \leq p < \infty$ , which are spaces of martingales with certain integrability properties, the most important being  $\mathcal{H}^1$ . We present first the  $\mathcal{H}^p$  and specialize then to p = 1 and p = 2. This is inspired by [6] and does explicitly not make use of the fundamental theorem of local martingales.

Main tool for the analysis are the Burkholder-Davis-Gundy inequalities:

**Theorem 6.1.** For every  $p \ge 0$  there are constants  $0 < c_p < C_p$  such that for every martingale

$$c_p E[[M,M]_{\infty}^{\frac{p}{2}}] \le E[(|M|_{\infty}^*)^p] \le C_p E[[M,M]_{\infty}^{\frac{p}{2}}]$$

holds true.

*Remark* 6.2. The inequalities follow from the same inequalities for discrete martingales, which can be proved by deterministic methods. In fact equations of the type

$$(h \bullet M)_T + [M, M]_T^{\frac{L}{2}} \le (|M|_T^*)^p \le C_p[M, M]_T^{\frac{L}{2}} + (g \bullet M)_T$$

hold true, with predictable integrands h, g and martingales M on a finite index set with upper bound T hold, see [1].

Let M be a martingale and let us take a sequence of refining partitions  $\Pi^n$  tending to identity, for which  $M^{\Pi^n} \to M$  in ucp and  $[M^{\Pi^n}, M^{\Pi^n}] \to [M, M]$  in ucp. Fix some time horizon T > 0, then by monotone convergence

$$E\left[\left(|M^{\Pi^n}|_T^*\right)^p\right] \to E\left[\left(|M|_T^*\right)^p\right]$$

as  $n \to \infty$ , since the sequence of partitions is refining. If  $E[(|M|_T^*)^p] = \infty$ , we obtain that all three quantities are infinity. If  $E[(|M|_T^*)^p] < \infty$  we obtain by dominated convergence that

$$E\left[\left[M^{\Pi^{m}} - M^{\Pi^{n}}, M^{\Pi^{m}} - M^{\Pi^{n}}\right]_{T}^{\frac{p}{2}}\right] \le E\left[\left(|M^{\Pi^{m}} - M^{\Pi^{n}}|_{T}^{*}\right)^{p}\right] \to 0$$

which means by

$$E\Big[\left|\left[M^{\Pi^{m}}, M^{\Pi^{m}}\right]^{\frac{1}{2}} - \left[M^{\Pi^{n}}, M^{\Pi^{n}}\right]_{T}^{\frac{1}{2}}\right|^{p}\Big] \le E\Big[\left[M^{\Pi^{m}} - M^{\Pi^{n}}, M^{\Pi^{m}} - M^{\Pi^{n}}\right]_{T}^{\frac{p}{2}}\Big]$$

the  $L^p$  convergence of the quadratic variations to [M, M]. This yields the result for any T > 0 and hence for  $T \to \infty$ .

Remark 6.3. The case p = 2 can be readily derived from Doob's maximal inequality, see Theorem 3.3, and we obtain

$$E\left[[M,M]_{\infty}\right] = E\left[M_{\infty}^{2}\right] \le E\left[\left(|M|_{\infty}^{*}\right)^{2}\right] \le 4E[M_{\infty}^{2}] = E\left[[M,M]_{\infty}\right].$$

**Definition 6.4.** Let  $p \ge 1$  be given. Define the vector space  $\mathcal{H}^p$  as set of martingales M where

$$\|M\|_{\mathcal{H}^p}^p := E\left[\left(|M|_{\infty}^*\right)^p\right] < \infty$$

holds true.

By the Burkholder-Davis-Gundy inequalities the following theorem easily follows: **Theorem 6.5.** For  $p \ge 1$  the space  $\mathcal{H}^p$  is a Banach space with equivalent norm

$$M \mapsto E\left[\left[M, M\right]_{\infty}^{\frac{p}{2}}\right]^{\frac{1}{p}}$$

For p = 2 the equivalent norm is in fact coming from a scalar product

 $(M, N) \mapsto E[[M, N]_T].$ 

Additionally we have the following continuity result:  $M^n \to M$  in  $\mathcal{H}^p$ , then  $(Y \bullet M_n) \to (Y \bullet M)$  in ucp for any left-continuous process  $Y \in \mathbb{L}$ . In particular  $[M^n, N] \to [M, N]$  in ucp.

In the next step we consider a weaker topology of  $L^p$  type on the set of simple predictable integrands. The following lemma tells about the closure with respect to this topology.

**Lemma 6.6.** Let A be an increasing finite variation process and V a predictable process with

$$(|V|^p \bullet A)_t < \infty$$

then there exists a sequence of bounded, simply predictable processes  $V^n$  in  $b\mathbb{S}$  such that

$$(|V - V^n|^p \bullet A)_t \to 0,$$

as  $n \to \infty$ .

*Proof.* By monotone class arguments it is sufficient to prove the lemma for LCRL processes, for which it is, however, clear, since they can be approximated in the ucp topology by simply predictable processes.  $\Box$ 

The main line of argument is to construct for predictable processes V satisfying certain integrability conditions with respect to [M, M] a stochastic integral  $(V \bullet M)$ for  $M \in \mathcal{H}^p$ . We take the largest space  $\mathcal{H}^1$  in order to stay as general as possible.

**Proposition 6.7.** Let  $M \in \mathcal{H}^1$  be fixed and let  $V^n$  be a sequence of bounded, simple predictable processes such that

$$E\left[\left(|V-V^n|^2 \bullet [M,M]\right)_{\infty}^{\frac{1}{2}}\right] \to 0$$

then the sequence  $(V^n \bullet M)$  is a Cauchy sequence in  $\mathcal{H}^1$  defining an element  $(V \bullet M)$ , which does only depend on V and not on the approximating sequence  $V^n$  and which is uniquely determined by

$$[(V \bullet M), N] = (V \bullet [M, N])$$

for martingales N.

*Proof.* This is a direct consequence of the Burkholder-Davis-Gundy inequalities, since

$$E\big[|(V^n \bullet M) - (V^m \bullet M)|_{\infty}^*)\big] \le C_1 E\big[((V^n - V^m)^2 \bullet [M, M])_{\infty}^{\frac{1}{2}}\big] \to 0$$

as  $n, m \to \infty$ . Whence  $(V \bullet M)$  is a well-defined element of  $\mathcal{H}^1$ , which only depends on V and not on the approximating sequence. For all martingales N and all simple predictable strategies the formula

$$[(V^n \bullet M), N] = (V^n \bullet [M, N])$$

holds true by basic rules for LCRL integrands. By passing to the limit we obtain the general result. Uniqueness is clear since [M, M] = 0 means M = by Burkholder-Davis-Gundy inequalities.

**Definition 6.8.** Let  $M \in \mathcal{H}^1$ , then we denote by  $L^1(M)$  the set of predictable processes V such that

$$E\left[\left(\left|V\right|^{2}\bullet\left[M,M\right]\right)_{\infty}^{\frac{1}{2}}\right]<\infty.$$

Apparently we have constructed a bounded linear map  $L^1(M) \to \mathcal{H}^1$ ,  $V \mapsto (V \bullet H)$ . The set of integrands  $L^1(M)$  is not the largest one, we can still generalize it by localization, which defines the set L(M): a predictable process V is said to belong to L(M) if  $(|V|^2 \bullet [M, M])^{\frac{1}{2}}$  is locally integrable, which means for bounded variation processes nothing else than just being finite, see [6]. Notice that this is the largest set of integrands given that we require that the integral is a semi-martingal having a quadratic variation, which coincides with  $(V^2 \bullet [M, M])$ . Notice also that by the same argument every local martingale is in fact locally  $\mathcal{H}^1$ , which in turn means that we can define for any semi-martingale a largest set of integrands.

### References

- Mathias Beiglböck and Pietro Siorpaes. Pathwise versions of the Burkholder-Davis-Gundy inequalities. preprint, 2013.
- [2] Klaus Bichteler. Stochastic integration and L<sup>p</sup>-theory of semimartingales. Ann. Probab., 9(1):49–89, 1981.
- [3] Olav Kallenberg. Foundations of modern probability. Probability and its Applications (New York). Springer-Verlag, New York, second edition, 2002.
- [4] Paul-André Meyer. Martingales and stochastic integrals. I. Lecture Notes in Mathematics. Vol. 284. Springer-Verlag, Berlin-New York, 1972.
- [5] Philip E. Protter. Stochastic integration and differential equations, volume 21 of Applications of Mathematics (New York). Springer-Verlag, Berlin, second edition, 2004. Stochastic Modelling and Applied Probability.
- [6] A. N. Shiryaev and A. S. Cherny. A vector stochastic integral and the fundamental theorem of asset pricing. *Tr. Mat. Inst. Steklova*, 237(Stokhast. Finans. Mat.):12–56, 2002.

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